

榮譽第一

國立東華大學
應用數學系

學年度第 學期

考試科目:

期中 期末 考試試卷

108



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01

1. Let $X \sim f(x; \theta)$, $\theta_0 \neq \theta_1$, $\theta_0, \theta_1 \in \Omega$, and given $d \in (0, 1)$
For testing $H_0: \theta = \theta_0$ v.s. $H_1: \theta = \theta_1$
i) (Existence)

\exists a test $\phi(X)$ and constant $c > 0$
s.t. i) $E_{\theta_0} \phi(X) = d$
ii) $\phi(X) = \begin{cases} 1, & \text{if } f(x; \theta_1) > c f(x; \theta_0) \\ 0, & \text{if } f(x; \theta_1) < c f(x; \theta_0) \end{cases}$

ii) (Sufficient condition)
If a test ϕ^* satisfies i) & ii) in i)
for some $c > 0$, then ϕ^* is MP test
within test whose level $= d$

i.e. $\forall \phi$ w/ $E_{\theta_0} \phi(X) = d$, $E_{\theta_1} \phi^*(X) \geq E_{\theta_1} \phi(X)$
vii) (Necessary condition)
If ϕ^* is a MP level d test
then ϕ^* satisfies i) & ii) in i) for some $c > 0$

2. X_1, \dots, X_n iid $\frac{\theta}{\gamma^2}$, $\theta \in \mathcal{X} \subset \mathcal{R}$, $0 < \theta_1 < \theta_0$, $d \in (0, 1)$
To test $H_0: \theta = \theta_0$ v.s. $H_1: \theta = \theta_1$

$f(x; \theta) = \frac{1}{\gamma^n} \exp\left(-\frac{x}{\gamma}\right)$
 $\phi(X) = \begin{cases} 1, & \text{if } f(x; \theta_1) > c f(x; \theta_0) \\ 0, & \text{if } f(x; \theta_1) < c f(x; \theta_0) \end{cases}$
i.e. UMP level d test rej H_0 iff $X_{(1)} > \frac{\theta_0}{\sqrt{d}}$
 \Rightarrow i) b's test is same as i) a's test

\exists a UMP level d test $\phi^*(X)$,
and $\phi_d = d$, $E_{\theta} \phi_d(X) = d$
 $\therefore \phi^*(X)$ is UMP
 $\therefore E_{\theta} \phi^*(X) \geq E_{\theta} \phi_d(X) = d$
 $\Rightarrow \phi^*$ is unbiased
i.e. ϕ^* is UMPU level d test

$d = E_{\theta_0} \phi(X) = P_{\theta_0}(X_{(1)} > k) = [P_{\theta_0}(X > k)]^n$, take $\gamma = \theta_0$
 $= [1 - P_{\theta_0}(X \leq k)]^n = \left[1 - \int_0^k \frac{\theta_0}{\gamma^2} dx\right]^n$
 $= \left[1 + \frac{\theta_0}{\gamma} \left| \frac{k}{\theta_0} \right.\right]^n = \left[1 + \frac{\theta_0}{k} - 1\right]^n = \left(\frac{\theta_0}{k}\right)^n$
 $\Rightarrow \frac{\theta_0}{k} = d^{\frac{1}{n}} \Rightarrow k = \frac{\theta_0}{\sqrt{d}}$

The power of ϕ
 $= E_{\theta_1} \phi(X) = P_{\theta_1}(X_{(1)} > \frac{\theta_0}{\sqrt{d}}) = [P_{\theta_1}(X > \frac{\theta_0}{\sqrt{d}})]^n$

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4. X_1, \dots, X_n iid $N(\theta, \sigma^2)$, $\sigma^2 > 0$

To test $H_0: \theta \geq \theta_0$ vs. $H_1: \theta < \theta_0$

Under $\omega_0 = \{ \theta = \theta_0 \}$

$$\hat{\theta}_{\omega_0} = \begin{cases} \bar{X}, & \text{if } \bar{X} \geq \theta_0 \\ \theta_0, & \text{if } \bar{X} < \theta_0 \end{cases}$$

Under $\Omega = \{ \theta = \theta \in \mathbb{R} \}$

$$\hat{\theta}_{\Omega} = \bar{X}$$

$$L(\theta; \omega_0; \mathbf{x}) = \frac{L(\theta; \omega_0; \mathbf{x})}{L(\theta; \Omega; \mathbf{x})} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum (X_i - \theta_{\omega_0})^2}{2\sigma^2}\right)}{\left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum (X_i - \bar{X})^2}{2\sigma^2}\right)}$$

$$= \begin{cases} 1, & \text{if } \bar{X} \geq \theta_0 \\ \frac{\exp\left(-\frac{\sum (X_i - \theta_0)^2}{2\sigma^2}\right)}{\exp\left(-\frac{\sum (X_i - \bar{X})^2}{2\sigma^2}\right)}, & \text{if } \bar{X} < \theta_0 \end{cases}$$

$< k$ w/ $k \in (0, 1)$

$$\Leftrightarrow \frac{\exp\left(-\frac{\sum X_i^2 - 2\theta_0 \sum X_i + n\theta_0^2}{2\sigma^2}\right)}{\exp\left(-\frac{\sum X_i^2 - 2\bar{X} \sum X_i + n\bar{X}^2}{2\sigma^2}\right)} < k, \bar{X} < \theta_0$$

$$\Leftrightarrow \exp\left(\frac{-\sum X_i + 2\theta_0 n\bar{X} - n\theta_0^2 + \sum X_i - 2n\bar{X} + n\bar{X}^2}{2\sigma^2}\right) < k$$

$$\Leftrightarrow \exp\left(\frac{-n\bar{X} + 2n\theta_0\bar{X} - n\theta_0^2}{2\sigma^2}\right) < k, \bar{X} < \theta_0$$

$$\Leftrightarrow \frac{-n(\bar{X} - 2\theta_0\bar{X} + \theta_0^2)}{2\sigma^2} < k', \bar{X} < \theta_0$$

$$\Leftrightarrow -(\bar{X} - \theta_0)^2 > k'', \bar{X} < \theta_0$$

$$\Leftrightarrow \bar{X} < c$$

To have level α

$$\alpha = P_{\theta_0}(\bar{X} < c) = P_{\theta_0}\left(\frac{\sqrt{n}(\bar{X} - \theta_0)}{\sigma} < \frac{\sqrt{n}(c - \theta_0)}{\sigma}\right)$$

$$\Rightarrow \frac{\sqrt{n}(c - \theta_0)}{\sigma} = z_{1-\alpha} \Rightarrow c = \theta_0 + \frac{\sigma}{\sqrt{n}} z_{1-\alpha}$$

i.e. a level α LRT test for H_0 iff $\bar{X} < \theta_0 + \frac{\sigma}{\sqrt{n}} z_{1-\alpha}$

$$f(\mathbf{x}; \theta) = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum (X_i - \theta)^2}{2\sigma^2}\right)$$

$$= \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum X_i^2}{2\sigma^2}\right) \exp\left(+\frac{2\theta \sum X_i}{2\sigma^2}\right) \exp\left(-\frac{n\theta^2}{2\sigma^2}\right)$$

$$= \underbrace{\left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum X_i^2}{2\sigma^2}\right)}_{h(\mathbf{x})} \underbrace{\exp\left(\frac{n\theta}{\sigma^2} \bar{X}\right)}_{T(\mathbf{x})} \underbrace{\exp\left(-\frac{n\theta^2}{2\sigma^2}\right)}_{h(\theta)}$$

w/ $h(\theta) = \frac{n\theta}{\sigma^2} \uparrow$ in θ

$\Rightarrow T$ has MLR in \bar{X}

UMP level α test for $H_0: \theta \geq \theta_0$ vs. $H_1: \theta < \theta_0$

$$\text{is } \phi(\mathbf{x}) = \begin{cases} 1, & \text{if } \bar{X} < k \\ \gamma, & = \\ 0, & > \end{cases}$$

w/ $k, \gamma \in [0, 1]$ s.t. $\alpha = E_{\theta_0} \phi(\mathbf{x})$

$$\alpha = E_{\theta_0} \phi(\mathbf{x}) = P_{\theta_0}(\bar{X} < k) = P_{\theta_0}\left(\frac{\sqrt{n}(\bar{X} - \theta_0)}{\sigma} < \frac{\sqrt{n}(k - \theta_0)}{\sigma}\right)$$

$$\Rightarrow \frac{\sqrt{n}(k - \theta_0)}{\sigma} = z_{1-\alpha} \Rightarrow k = \theta_0 + \frac{\sigma}{\sqrt{n}} z_{1-\alpha}$$

i.e. UMP level α test for H_0 iff $\bar{X} < \theta_0 + \frac{\sigma}{\sqrt{n}} z_{1-\alpha}$

Hence, LRT is UMP

$$1) P_{\theta}(\bar{X} > \theta_0 + \frac{\sigma}{\sqrt{n}} z_{1-\alpha}) \leq \beta$$

$$= P_{\theta}\left(\frac{\sqrt{n}(\bar{X} - \theta)}{\sigma} > \frac{\sqrt{n}(\theta_0 - \theta + \frac{\sigma}{\sqrt{n}} z_{1-\alpha})}{\sigma}\right) \leq \beta$$

$$\Rightarrow \frac{\sqrt{n}(\theta_0 - \theta) + z_{1-\alpha}}{\sigma} \geq z_{\beta} \quad \text{w/ } \alpha = 0.05, \beta = 0.2, \theta = \theta_0 - 0.1\sigma$$

$$\frac{\sqrt{n}(\theta_0 - \theta_0 + 0.1\sigma)}{\sigma} \geq z_{0.2} - z_{0.95}$$

$$0.1\sqrt{n} \geq 0.845 - (-1.645) = 2.49$$

$$\sqrt{n} \geq 24.9$$

$$n \geq 6201$$

\Rightarrow Take $n = 6201$

5. X_1, \dots, X_n iid $0 \leq (1-x)^{\theta-1}$, $x \in (0, 1)$, $\theta > 0$, $\theta = \text{unknown}$

$$f(x; \theta) = \theta^n \prod_{i=1}^n (1-x_i)^{\theta-1} = \theta^n \prod_{i=1}^n \exp[(\theta-1) \log(1-x_i)]$$

$$= \theta^n \exp[(\theta-1) \sum \log(1-x_i)]$$

$$\underbrace{\theta^n}_{h(\theta)} \underbrace{\exp[(\theta-1) \sum \log(1-x_i)]}_{T(\mathbf{x})} \underbrace{1}_{h(\mathbf{x})}$$

w/ $h(\theta) = \theta^n \uparrow$ in θ

$\Rightarrow T$ has MLR in $\sum \log(1-x_i)$

note Let $Y = -\log(1-X)$, then $X = 1 - e^{-Y}$

$$f_Y(y) = f_X(1 - e^{-y}) |e^{-y}|$$

$$= \theta (1 - (1 - e^{-y}))^{\theta-1} \cdot e^{-y} = \theta e^{-y(\theta-1)} e^{-y}$$

$$= \theta e^{-\theta y}, \theta > 0, y > 0$$

$$Y \sim \text{Gamma}(1, \theta), \sum Y_i \sim \text{Gamma}(n, \theta)$$

$$f_W(w) = \frac{1}{\Gamma(n)(\theta)^n} w^{n-1} \exp\left(-\frac{w}{\theta}\right) = \frac{m}{z} \Rightarrow m = z\theta w$$

$$\geq \theta \sum Y_i = -\theta \sum \log(1-X) \sim \text{Gamma}(n, \theta) = \chi^2_{2n}$$

UMP level α test for $H_0: \theta \leq \theta_0$ vs. $H_1: \theta > \theta_0$

$$\text{is } \phi(\mathbf{x}) = \begin{cases} 1, & \text{if } \sum \log(1-x_i) > k \\ \gamma, & = \\ 0, & < \end{cases}$$

w/ $k, \gamma \in [0, 1]$ s.t. $\alpha = E_{\theta_0} \phi(\mathbf{x})$



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$d = E_{\theta_0} \psi(\mathbf{X}) = P_{\theta_0}(\sum \log(1-X_i) > k)$, take $\nu = 0$	$(\text{case 1, } \theta > \theta_0)$
$= P_{\theta_0}(-\sum \log(1-X_i) < -\sum \log(1-X_i) > k)$	$E_{\theta} \psi(\mathbf{X}) = P_{\theta}(X_{(n)} > \theta_0) + P_{\theta}(X_{(n)} < \theta_0 d^{\frac{1}{n}})$
$\Rightarrow -\sum \log(1-X_i) < -\sum \log(1-X_i) > k$	$= 1 - P_{\theta}(X_{(n)} \leq \theta_0) + [P_{\theta}(X < \theta_0 d^{\frac{1}{n}})]^n$
$\Rightarrow -2\theta_0 k = \chi_{2n, 1-\alpha}^2 \Rightarrow k = -\frac{1}{2\theta_0} \chi_{2n, 1-\alpha}^2$	$= 1 - [P_{\theta}(X < \theta_0)]^n + \left[\int_0^{\theta_0 d^{\frac{1}{n}}} \frac{1}{\theta} dx \right]^n$
i.e. UMP level α test w.r.t. H_0	$= 1 - \left[\int_0^{\theta_0} \frac{1}{\theta} dx \right]^n + \left(\frac{\theta_0 d^{\frac{1}{n}}}{\theta} \right)^n$
$\text{iff } \sum \log(1-X_i) > -\frac{1}{2\theta_0} \chi_{2n, 1-\alpha}^2$	$= 1 - (1-\alpha) \left(\frac{\theta_0}{\theta} \right)^n$
(b) The power of $\psi = E_{\theta} \psi(\mathbf{X}) = P_{\theta}(\sum \log(1-X_i) > k)$	$(\text{case 2, } \theta < \theta_0)$
$= P_{\theta}(-\sum \log(1-X_i) < -\sum \log(1-X_i) > k)$	$E_{\theta} \psi(\mathbf{X}) = P_{\theta}(X_{(n)} > \theta_0) + P_{\theta}(X_{(n)} < \theta_0 d^{\frac{1}{n}})$
$\Rightarrow \frac{\theta}{\theta_0} \chi_{2n, 1-\alpha}^2 \geq \chi_{2n, 1-\beta}^2$ w/ $d=0.05, \theta_1=1000-\theta$ $\beta=0.9, \theta_0=750$	$= [P_{\theta}(X < \theta_0 d^{\frac{1}{n}})]^n$
$4 \chi_{2n, 0.95}^2 \geq \chi_{2n, 0.1}^2$	$= \left[\int_0^{\theta_0 d^{\frac{1}{n}}} \frac{1}{\theta} dx \right]^n$
\Rightarrow Take n be the smallest integer s.t. \ast holds	$= \left(\frac{\theta_0 d^{\frac{1}{n}}}{\theta} \right)^n$
X_1, \dots, X_n iid $U(0, \theta), \theta > 0$	$= d \left(\frac{\theta_0}{\theta} \right)^n$
$\theta_2 > \theta_1, \frac{f(x; \theta_2)}{f(x; \theta_1)} = \frac{\left(\frac{1}{\theta_2}\right)^n \mathbb{I}(0 \leq X_i \leq \theta_2)}{\left(\frac{1}{\theta_1}\right)^n \mathbb{I}(0 \leq X_i \leq \theta_1)}$	$\text{UMP level } \alpha \text{ test for } H_0: \theta \leq \theta_0 \text{ vs. } H_1: \theta > \theta_0$
$= \left(\frac{\theta_1}{\theta_2}\right)^n \mathbb{I}(X_{(n)} \leq \theta_1)$	$\text{is } \psi_1(\mathbf{X}) = \begin{cases} 1, & \text{if } X_{(n)} > k \\ \gamma, & \text{if } X_{(n)} = k \\ 0, & \text{if } X_{(n)} < k \end{cases}$
$= \begin{cases} \left(\frac{\theta_1}{\theta_2}\right)^n, & \text{if } 0 \leq X_{(n)} \leq \theta_1 \\ 0, & \text{if } X_{(n)} > \theta_1 \end{cases}$	$\text{w/ } k, \gamma \in [0, 1] \text{ s.t. } d = E_{\theta_0} \psi_1(\mathbf{X})$
$\Rightarrow \frac{f(x; \theta_2)}{f(x; \theta_1)} \geq 1$ in $X_{(n)}$	$d = E_{\theta_0} \psi_1(\mathbf{X}) = P_{\theta_0}(X_{(n)} > k) = 1 - P_{\theta_0}(X_{(n)} \leq k)$
$\Rightarrow \psi$ has MLR in $X_{(n)}$	$= 1 - [P_{\theta_0}(X \leq k)]^n = 1 - \left[\int_0^k \frac{1}{\theta_0} dx \right]^n$
To test $H_0: \theta > \theta_0$ vs. $H_1: \theta \neq \theta_0$	$= 1 - \left(\frac{k}{\theta_0} \right)^n$
$\psi(\mathbf{X}) = \begin{cases} 1, & \text{if } X_{(n)} > \theta_0 \text{ or } X_{(n)} < k \\ 0, & \text{v.w.} \end{cases}$	$\Rightarrow \left(\frac{k}{\theta_0} \right)^n = 1 - \alpha \Rightarrow k = \theta_0 (1 - \alpha)^{\frac{1}{n}}$
$\text{w/ } k, \text{ s.t. } d = E_{\theta_0} \psi(\mathbf{X})$	$\text{i.e. UMP level } \alpha \text{ test w.r.t. } H_0 \text{ iff } X_{(n)} > \theta_0 (1 - \alpha)^{\frac{1}{n}}$
$d = E_{\theta_0} \psi(\mathbf{X}) = P_{\theta_0}(X_{(n)} > \theta_0) + P_{\theta_0}(X_{(n)} < k)$	$\text{The power of } \psi_1 = E_{\theta} \psi_1(\mathbf{X}) = P_{\theta}(X_{(n)} > \theta_0 (1 - \alpha)^{\frac{1}{n}})$
$= [P_{\theta_0}(X < k)]^n = \left[\int_0^k \frac{1}{\theta_0} dx \right]^n = \left(\frac{k}{\theta_0} \right)^n$	$= 1 - P_{\theta}(X_{(n)} \leq \theta_0 (1 - \alpha)^{\frac{1}{n}}) = 1 - [P_{\theta}(X \leq \theta_0 (1 - \alpha)^{\frac{1}{n}})]^n$
$\Rightarrow d^{\frac{1}{n}} = \frac{k}{\theta_0} \Rightarrow k = \theta_0 d^{\frac{1}{n}}$	$= 1 - \left[\int_0^{\theta_0 (1 - \alpha)^{\frac{1}{n}}} \frac{1}{\theta} dx \right]^n = 1 - \left(\frac{\theta_0 (1 - \alpha)^{\frac{1}{n}}}{\theta} \right)^n$
$\text{i.e. } \psi(\mathbf{X}) = \begin{cases} 1, & \text{if } X_{(n)} > \theta_0 \text{ or } X_{(n)} < \theta_0 d^{\frac{1}{n}} \\ 0, & \text{v.w.} \end{cases}$	$= 1 - (1 - \alpha) \left(\frac{\theta_0}{\theta} \right)^n, \therefore \psi_1 \text{ is UMP}$
$\text{The power of } \psi = E_{\theta} \psi(\mathbf{X}) = \begin{cases} E_{\theta} \psi(\mathbf{X}), & \theta > \theta_0 \\ E_{\theta} \psi^*(\mathbf{X}), & \theta < \theta_0 \end{cases}$	$\geq E_{\theta} \psi^*(\mathbf{X}), \forall \psi^* \text{ is a level } \alpha \text{ test for } H_1$

UMP level α test for $H_0: \theta \geq \theta_0$ vs. $H_1: \theta < \theta_0$

$$\phi(x) = \begin{cases} 1, & \text{if } X_{(n)} < k \\ 0, & \text{if } X_{(n)} \geq k \end{cases}$$

$$\forall k, \forall \theta \in [0, 1] \text{ s.t. } \alpha = E_{\theta_0} \phi(X)$$

$$\begin{aligned} \alpha &= E_{\theta_0} \phi(X) = P_{\theta_0}(X_{(n)} < k) = [P_{\theta_0}(X < k)]^n \\ &= \left[\int_0^k \frac{1}{\theta_0} d\tau \right]^n = \left(\frac{k}{\theta_0} \right)^n \\ \Rightarrow k &= \theta_0 \alpha^{\frac{1}{n}} \end{aligned}$$

i.e. UMP level α test $\phi(x)$ for H_0 if $X_{(n)} < \theta_0 \alpha^{\frac{1}{n}}$

The power of $\phi(x) = E_{\theta} \phi(x) = P_{\theta}(X_{(n)} < \theta_0 \alpha^{\frac{1}{n}})$

$$\begin{aligned} &= [P_{\theta}(X < \theta_0 \alpha^{\frac{1}{n}})]^n = \left[\int_0^{\theta_0 \alpha^{\frac{1}{n}}} \frac{1}{\theta} d\tau \right]^n \\ &= \left(\frac{\theta_0 \alpha^{\frac{1}{n}}}{\theta} \right)^n = \alpha \left(\frac{\theta_0}{\theta} \right)^n, \quad \therefore \phi(x) \text{ is UMP} \end{aligned}$$

$\Rightarrow E_{\theta} \phi^*(X)$, $\forall \phi^*$ is a level α test for H_1

$$\Rightarrow E_{\theta} \phi(X) = \begin{cases} (1-\alpha) \left(\frac{\theta_0}{\theta} \right)^n, & \text{if } \theta \geq \theta_0 \\ \alpha \left(\frac{\theta_0}{\theta} \right)^n, & \text{if } \theta < \theta_0 \end{cases}$$

$$= \begin{cases} E_{\theta} \phi_1(X) \\ E_{\theta} \phi_2(X) \end{cases}$$

$$\Rightarrow E_{\theta} \phi^*(X),$$

$\forall \phi^*$ is a level α test for H_1

Hence, the test $\phi(X)$ is a UMP level α

test for testing $H_0: \theta = \theta_0$ vs. $H_1: \theta \neq \theta_0$.

1.	A \ B	Yes	No	
	Yes	10	30	40
	No	40	20	60
		50	50	100

Let A: treatment, B: control

To test $H_0: A \perp B$ vs. $H_1: A \not\perp B$

$$\begin{aligned} E &= \begin{bmatrix} E_{00} & E_{01} \\ E_{10} & E_{11} \end{bmatrix} = \begin{bmatrix} \frac{40 \times 50}{100} & \frac{40 \times 50}{100} \\ \frac{60 \times 50}{100} & \frac{60 \times 50}{100} \end{bmatrix} \\ &= \begin{bmatrix} 20 & 20 \\ 30 & 30 \end{bmatrix} \end{aligned}$$

$$\chi^2 = \frac{(10-20)^2}{20} + \frac{(30-20)^2}{20} + \frac{(40-30)^2}{30} + \frac{(20-30)^2}{30}$$

$$= 5 + 5 + \frac{10}{3} + \frac{10}{3} = 16.\bar{6}$$

$$\chi^2 = 16.\bar{6} > \chi^2_{(2-1)(2-1), 0.05} = \chi^2_{1, 0.05} = 3.84$$

\Rightarrow reject H_0 i.e. ...

8. To test $H_0: X \sim \text{Gamma}(0, 1), 1 \leq \theta \leq 6$

vs. $H_1: Y \sim P(1), 1 \leq \theta \leq 6$

∴ No mind.

3.84

0

0.5

0.5

0

10