Here is sketches of proofs for Homework problems #2, #3 in §3.

EX 1 (§3.8 #2) Let X be a random variable with discrete uniform pmf on $\{\pm 3, \pm 2, \pm 1, 0\}$. Let g be pmf of the random variable $Y = r(X) = X^2 - X$. Calculate the pmf g for Y. <u>Solution</u>. Note that for this question, it might be easier to follow the definition (1) then trying to figure out the inverse of r over various subdomains.

First we will make clear what the values that Y can take. Specifically, let

$$X = \{x/f(x) > 0\} = \{\pm 3, \pm 2, \pm 1, 0\}$$

We would like to get

$$Y = \{y|g(y) > 0\} = \{y|y = r(x), x \in X\}.$$

Note

Х	-3	-2	-1	0	1	2	3
У	12	6	2				<u>.</u>

where 1[A] is the indicator function of event A. Since

$$r(x) = x(2 - x) \quad y \quad x^2 - 2x + y \quad 0$$

(x 1 + 1 - y) or (x 1 - 1 - y).

Note for y = r(x) = x(2 - x), the derivative y = -(x - 1). Thus for x = [0, 2] (including 0, 2 makes no di erences since f is a pdf), y = r(x) is increasing for x < 1; decreasing for x > 1. In addition, it reaches its maximum 1 when x = 1 and attains its minimum 0 at x = 0, 2. Hence G(y) = 0 if y < 0; = 1 if y > 1; for y = [0, 1],

$$(2) = {2 \atop 0} 1[(x \quad 1 + \overline{1 - y}) \text{ or } (x \quad 1 - \overline{1 - y})] \quad \frac{x}{2} \, dx \\ = {2 \atop 1 + \overline{1 - y}} + {1 - \overline{1 - y} \atop 2} \frac{x}{2} \, dx = \frac{x^2}{4} {2 \atop 1 + \overline{1 - y}} + \frac{x^2}{4} {1 - \overline{1 - y}} \\ = 1 - \frac{1}{4} [(1 + \overline{1 - y})^2 - (1 - \overline{1 - y})^2] = 1 - \overline{1 - y}.$$

To summarize,

$$G(y) = \begin{array}{ccc} 0, & y < 0 \\ 1 - \overline{1 - y}, & y \quad [0, 1] \\ 1, & y > 1. \end{array}$$

Thus the pdf

$$g(y) = \frac{d}{dy}G(y) = \frac{1}{2(1-y)}, \quad y = (0, 1), \\ 0, \qquad \text{Otherwise}.$$

Remark 1 A way to check whether we do get the right pdf/pmf is to calculate the expectation of Y (if exists) in two ways and confirm the equivalence. Specifically, for the case when X has pdf f and Y = r(X) has pdf g

$$E(Y) = yg(y) dy$$
(3)

$$E(r(X)) = r(x)f(x) dx.$$
(4)

*§*3. : #3 as an illustration.

$$E(r(X)) = EX(2 - X) = \int_{0}^{2} 2x - x^{2} dx = \frac{2}{3}.$$

$$E(Y) = \int_{0}^{1} y \frac{1}{2 \cdot 1 - y} dy, \quad \text{let } y = \sin^{2}, \quad (0, /2)$$

$$= \int_{0}^{/2} \frac{\sin^{2}}{2 \cdot 1 - \sin^{2}} 2\sin \cos d$$

$$= \int_{0}^{/2} \sin^{3} d = \int_{0}^{2} (\cos^{2} - 1) d\cos d$$

$$= \frac{z^{3}}{3} - z \int_{1}^{0} = \frac{2}{3}.$$

This confirms that we did get the right pdf! The same idea works for moment generating functions as well. It might be easier to calculate in some cases.

Finally, I supplement a proof of a claim made in the Textbook. This proof is for the curious enjoyment only. :)

EX 3 (Example 4.4.1, P 204)

$$- \frac{|x|^{k} e_{-(x-3)^{2}}}{k} dx < k N.$$

Proof. For any positive integer *k*, by L'Hôpital's Rule

Thus for any > 0, there exists an M() > 0 such that

$$|X|^{k} \underline{e}_{(x-3)^{2}} < , \qquad |X| > M.$$

Therefore

$$- \frac{|x|^{k}e^{(x-3)^{2}} dx = 2}{0} \frac{x^{k}e^{(x-3)^{2}} dx}{2} = 2(\frac{M}{0} + \frac{M}{M})x^{k}\frac{e}{(x-3)^{2}} dx$$

$$- \frac{M}{2} \frac{x^{k}e^{(x-3)^{2}} dx}{2M^{k} + (x-3)^{2}} \frac{dx}{dx}$$

You can refer to P 277, Example 1, Mathematical Analysis, 2nd Edition, Apostol for more